

1 Set Theory and Intersection

1.1 Union and Intersection

A union B = {x: x in A or x in B}

A intersection B = {x: x in A and x in B}

1.2 Complement

A complement = {x: x not in A}

1.3 Disjoint Sets

Two sets A\_i and A\_j are disjoint if

A\_i intersection A\_j = empty set for all i, j not equal

1.4 Collectively Exhaustive Sets

Sets A\_1, ..., A\_n are collectively exhaustive if

Union from i=1 to n of A\_i = S

1.5 Partition

Sets A\_1, ..., A\_n are called a partition of S if A\_i, ..., A\_n are disjoint and collectively exhaustive.

1.6 Properties of Sets

1.6.1 Commutative

A intersection B = B intersection A, A union B = B union A

1.6.2 Associative

A union (B union C) = (A union B) union C

A intersection (B intersection C) = (A intersection B) intersection C

1.6.3 Distributive

A union (B intersection C) = (A union B) intersection (A union C)

A intersection (B union C) = (A intersection B) union (A intersection C)

1.7 Relative Complement/Difference

A - B = {x: x in A and x not in B}

A - B = A intersection B complement

1.7.1 DeMorgan's

(A union B) complement = A complement intersection B complement, (A intersection B) complement = A complement union B complement

2 Probability Theory Introduction

2.1 Relative Frequency

Suppose that an experiment is repeated n times under identical conditions. Let N\_0(n), N\_1(n), ..., N\_k(n) be the number of times the outcome k happens. Then the relative frequency of outcome k is

f\_k(n) = N\_k(n)/n, where limit as n approaches infinity of f\_k(n) = p\_k

2.2 Axioms of Probability

P(A) >= 0, P(S) = 1

A intersection B = empty set implies P(A union B) = P(A) + P(B)

P(A union B) = P(A) + P(B) - P(A intersection B)

If A\_1, A\_2 is a sequence of events s.t. A\_i intersection A\_j = empty set for i not equal j

2.3 Bayesian Probability

P(A union B) = P(A) + P(B) - P(A intersection B)

3 Counting Methods and Sampling

Table with 2 columns: Description and Value. Includes permutations, ordered samples with/without replacement, and unordered samples with/without replacement.

3.1 Binomial Coefficient

n choose k = n! / (k! \* (n-k)!), n choose n-k = n choose k

4 Conditional Probability

If A and B are related, then the conditional probability of A given that B (and P[B] > 0) has occurred is

P[A|B] = P[A intersection B] / P[B]

5 Theorem of Total Probability

For B\_1, B\_2, ..., B\_n mutually exclusive events whose union equals the sample space S (e.g. B\_1, ..., B\_n is a partition of S), then

P[A] = P[A|B\_1]P[B\_1] + ... + P[A|B\_n]P[B\_n]

6 Bayes Rule

For B\_1, B\_2, ..., B\_n a partition of sample space S,

P[B\_j|A] = P[A intersection B\_j] / P[A] = P[A|B\_j]P[B\_j] / (sum from k=1 to n of P[A|B\_k]P[B\_k])

7 Independence

If knowledge of the occurrence of event B does not alter the probability of event A, then event A is independent of B.

P[A] = P[A|B] = P[A intersection B] / P[B]

Define A, B to be independent if

P[A intersection B] = P[A]P[B]

Then

P[A|B] = P[A], P[B|A] = P[B]

P[A complement intersection B complement] = P[A complement]P[B complement]

7.1 Notes on Independence

If two events have nonzero probability (P[A] > 0, P[B] > 0), and are mutually exclusive, then they cannot be independent.

7.2 Triplet Independence

For three events A, B, C to be independent,

- A, B, C Pairwise Independent
knowledge of occurrence of any two events (e.g. A, B) should not effect the prob. of the third (C)

7.2.1 Pairwise Independence

P[A intersection B] = P[A]P[B], P[A intersection C] = P[A]P[C]

P[B intersection C] = P[B]P[C]

7.2.2 Independence of Events

P[C|A intersection B] = P[A intersection B intersection C] / P[A intersection B] = P[C]

Finally, for Triplet Independence, we must have

P[A intersection B intersection C] = P[A]P[B]P[C]

8 Sequential Experiments

8.1 Bernoulli Trials

Let k be the num of successes in n independent Bernoulli trials. Then the probabilities of k are given by binomial probability law

P\_n(k) = (n choose k) p^k (1-p)^(n-k) for k = 0, ..., n

8.2 Multinomial Probability Law

Let B\_1, B\_2, ..., B\_M be a partition of the sample space S, and let P[B\_j] = p\_j. Also, the events are disjoint:

p\_1 + p\_2 + ... + p\_M = 1

The multinomial probability law is

P((k\_1, ..., k\_M)) = n! / (k\_1! \* ... \* k\_M!) \* p\_1^k\_1 \* ... \* p\_M^k\_M

8.3 Geometric Probability Law

The probability that more than K trials are required before a success (with probability p, q = 1 - p) occurs in a series of repeated independent Bernoulli trials is

P[m > K] = p \* sum from m=K+1 to infinity of q^(m-1) = pq^K \* 1/(1-q) = q^K

The probability that K trials are required for a success (with probability p, q = 1 - p) is

P[m = K] = (p)(1-p)^(K-1) = pq^(K-1)

8.3.1 Hypergeometric Distribution

P(X = k) = (K choose k) \* (N-K choose n-k) / (N choose n)

Where K is the number of success in the population, k is the number of observed successes, N is the population size, and n is the sample size.

9 Random Variables (RV)

A Random Variable X is a function that assigns a real number X(C\_i) to each outcome C\_i in the sample space of a random experiment.

10 Discrete Random Variable (DRV)

A Discrete Random Variable X is defined as a random variable that assumes values from a countable set.

10.1 PMF

P\_X(x) = P[X = x] = P[{C: X(C) = x}] x in R

For x\_k in S\_X, P\_X(x\_k) = P[A\_k]

10.1.1 PMF Properties

P\_X(x) >= 0 for all x

sum over x in S\_X of P\_X(x) = sum over k of P\_X(x\_k) = sum over k of P[A\_k] = 1

P[X in B] = sum over x in B of P\_X(x) where B subset S\_X

10.2 Conditional PMF

Let X be a DRV with PMF P\_X(x), and exists C, P[C] > 0. The Conditional PMF is given by

P\_X(x|C) = P[X = x|C] = P[(X = x) intersection C] / P[C]

10.3 Expected Value

The expected value (or mean) of a DRV is

E[X] = sum over x in S\_X of x P\_X(x) = sum over k of x\_k P\_X(x\_k)

exists E[|x|] = sum over k of |x\_k| P\_X(x\_k) < infinity

10.4 Variance, Standard Deviation

The variance of a random variable X is

sigma\_X^2 = VAR[X] = E[(X - E[X])^2] = E[X^2] - (E[X])^2

The Standard Deviation is

sigma\_X = STD[X] = sqrt(VAR[X])

10.5 Expected Value and Variance Properties

E[g(X) + h(X)] = E[g(X)] + E[h(X)]

E[aX] = aE[X], E[X + c] = E[X] + c

VAR(cX) = c^2 VAR(X), VAR(X + c) = VAR(X)

10.6 Conditional Expected Value

For X a DRV, and suppose we know B has occurred,

m\_X|B = E[X|B] = sum over x in S\_X of x P\_X(x|B)

= sum over k of x\_k P\_X(x\_k|B)

10.7 Conditional Variance

VAR[X|B] = E[(X - m\_X|B)^2|B]

sum over k=1 to infinity of (X\_k - m\_X|B)^2 P\_X(x\_k|B) = E[X^2|B] - m\_X^2|B

11 Cumulative Distribution Function

PMF's use events {X = b}, whereas Cumulative Distribution Functions (CDF) use events {X <= b}.

F\_X(x) = P[X <= x]

0 <= F\_X(x) <= 1

lim as x approaches -infinity of F\_X(x) = 0, lim as x approaches infinity of F\_X(x) = 1

F\_X(a) <= F\_X(b) for all a < b

F\_X(b) = lim as h approaches 0 of F\_X(b+h) = F\_X(b^+)

P[a < X <= b] = F\_X(b) - F\_X(a)

P[X = b] = F\_X(b) - F\_X(b^-)

P[X > x] = 1 - F\_X(x)

11.2 CDF of a Discrete RV

F\_X(x) = sum over x\_k <= x of P\_X(x\_k) = sum over k of P\_X(x\_k) u(x - x\_k)

11.3 CDF of a Continuous RV

F\_X(x) = integral from -infinity to x of f(t) dt

11.4 Conditional CDF

F\_X(x|C) = P[(X <= x) intersection C] / P[C] if P[C] > 0

12 Probability Density Function

f\_X(x) = d/dx F\_X(x)

12.1 Properties of the PDF

f\_X(x) >= 0, 1 = integral from -infinity to infinity of f\_X(x) dx

P[a <= X <= b] = integral from a to b of f\_X(x) dx

F\_X(x) = integral from -infinity to x of f\_X(t) dt

12.2 PDF of a Discrete RV

u(x) = integral from -infinity to x of delta(t) dt

f\_X(x) = d/dx F\_X(x) = sum over k of P\_X(x\_k) delta(x - x\_k)

12.3 Conditional PDF

f\_X(x|C) = d/dx F\_X(x|C)

12.4 Application of Theorem of Total Probability

Suppose events B\_1, B\_2, ..., B\_n partition the sample space S.

F\_X(x) = sum from i=1 to n of P[X <= x|B\_i] P[B\_i]

= sum from i=1 to n of F\_X(x|B\_i) P[B\_i]

f\_X(x) = d/dx F\_X(x) = sum from i=1 to n of f\_X(x|B\_i) P[B\_i]

13 Gaussian (Normal) RV

The PDF for the Gaussian Random Variable is given in the table.

13.1 Gaussian CDF

phi is the CDF for a standard Gaussian.

phi(z) = P((X - m)/sigma <= z) = P[X <= x] = F\_X(x)

phi(x) = 1/sqrt(2pi) integral from -infinity to x of e^(-t^2/2) dt

13.2 Q Function

Q(x) = 1/sqrt(2pi) integral from x to infinity of e^(-t^2/2) dt

Q(z) = 1 - phi(z) = P[X > x]

Q(0) = 1/2, Q(-x) = 1 - Q(x)

13.3 Standard Gaussian RV

To move from any Gaussian to Standard (i.e. X ~ N(m, sigma^2) -> z ~ N(0, 1)), use

z = (x - m)/sigma

14 Other Features of CRV's

14.1 Expected Value

E[X] = integral from -infinity to infinity of t f\_X(t) dt

14.1.1 Expected Value of Y=g(X)

E[Y] = integral from -infinity to infinity of g(x) f\_X(x) dx

E[X^n] = (-1)^n d^n/ds^n X(s)|\_s=0

14.1.2 Conditional Expected Value

E[X|A] = integral from -infinity to infinity of x f\_X(x|A) dx

14.2 Variance, Standard Deviation

The variance of a random variable X is

VAR[X] = E[(X - E[X])^2] = E[X^2] - (E[X])^2

The standard deviation is

STD[X] = sqrt(VAR[X])

14.3 Nth Moment

The nth moment of a random variable X is given by

E[X^n] = integral from -infinity to infinity of x^n f\_X(x) dx

15 Functions of RVs - CDF, PDF of Y

f\_Y(y) = sum from i=1 to n of f\_X(x\_i) / |g'(x\_i)|

f\_Y(y) = sum over x of f\_X(x) |dx/dy| = sum over k of f\_X(x\_k) |dx/dy| = sum over k of f\_X(x\_k) |dx/dy| = sum over k of f\_X(x\_k) |dx/dy|

16 Bounds on Probability

16.1 Markov Inequality

Suppose X is a RV with mean E[X]. Then

P[X >= a] <= E[X]/a for X nonnegative

16.2 Chebyshev Inequality

Suppose X is a RV with mean m = E[X] and variance sigma^2.

P[|X - m| >= a] <= sigma^2/a^2, D^2 = (X - m)^2 ->

P[D^2 >= a^2] <= E[(X - m)^2] / a^2 = sigma^2/a^2

16.3 Chernoff Bound

P[X <= a] = e^(-sa) E[e^sX]

17 Characteristic Function

phi\_X(w) = E[e^jwX] = integral from -infinity to infinity of f\_X(x) e^jwX dx

f\_X(x) = 1/(2pi) integral from -infinity to infinity of phi\_X(w) e^-jwX dw

17.1 Characteristic Function for DRV's

phi\_X(w) = sum over k of P\_X(x\_k) e^jwX\_k, X a DRV

phi\_X(w) = sum from -infinity to infinity of P\_X(k) e^jwk, X in Z

17.2 Moment Theorem

E[X^n] = 1/j^n d^n/dw^n phi\_X(w)|\_w=0

18 Moment Generating Function

M(s) = E[e^sX] = Phi(-js)

19 Probability Generating Function

G\_N(z) = E[z^N] = sum from k=0 to infinity of p\_N(k) z^k

19.1 Characteristic Function

G\_N(e^jw) = phi\_N(w)

19.2 PMF Relationship

PMF: P\_N(k) = 1/k! d^k/dz^k G\_N(z)|\_z=0

20 Laplace Transform of PDF

X(s) = integral from 0 to infinity of f\_X(x) e^(-sx) dx = E[e^(-sX)]

21 Joint PMF

P\_{X,Y}(x,y) = P[{X=x} intersection {Y=y}]

P[X in B] = sum over (x\_j, y\_k) in B of P\_{X,Y}(x\_j, y\_k)

sum from j=1 to infinity sum from k=1 to infinity of P\_{X,Y}(x\_j, y\_k) = 1

22 Marginal PMF

P\_X(x\_j) = P[X = x\_j] = sum from k=1 to infinity of P\_{X,Y}(x\_j, y\_k)

23 Joint CDF

F\_{X,Y}(x\_1, y\_1) = P[X <= x\_1, Y <= y\_1]

23.1 Properties of the Joint CDF

F\_{X,Y}(x\_1, y\_1) <= F\_{X,Y}(x\_2, y\_2)

for x\_1 <= x\_2, y\_1 <= y\_2

F\_{X,Y}(x\_1, -infinity) = 0, F\_{X,Y}(-infinity, y\_1) = 0, F\_{X,Y}(infinity, infinity) = 1

F\_X(x\_1) = F\_{X,Y}(x\_1, infinity), F\_Y(y\_1) = F\_{X,Y}(infinity, y\_1)

lim as x approaches a+ of F\_{X,Y}(x, y) = F\_{X,Y}(a, y)

lim as x approaches b+ of F\_{X,Y}(x, y) = F\_{X,Y}(b, y)

P[x\_1 < X < x\_2, y\_1 < Y < y\_2] = F\_{X,Y}(x\_2, y\_2) - F\_{X,Y}(x\_1, y\_2) - F\_{X,Y}(x\_2, y\_1) + F\_{X,Y}(x\_1, y\_1)

24 Joint PDF

f\_{X,Y}(x,y) = d^2 F\_{X,Y}(x,y) / (dx dy)

P[X in B] = integral over B of f\_{X,Y}(x,y) dx dy

F\_{X,Y}(x,y) = P[X <= x, Y <= y]

F\_{X,Y}(x,y) = integral from -infinity to y integral from -infinity to x of f\_{X,Y}(

$$E[X + Y] = E[X] + E[Y]$$

**27.1 Expected Value and Independence**  
Let  $g(X, Y) = g_1(X)g_2(Y)$ , and  $X, Y$  independent

$$Z = XY \leftrightarrow E[Z] = E[XY] = E[X]E[Y]$$

$$E[g(X, Y)] = E[g_1(X)]E[g_2(Y)]$$

**28 Joint Moment**  
If  $X, Y$  discrete:

$$E[X^j Y^k] = \sum_i \sum_n x_i^j y_n^k p_{XY}(x_i, y_n)$$

If  $X, Y$  jointly continuous:

$$E[X] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x^j y^k f_{XY}(x, y) dx dy$$

**28.1 Correlation**

$$E[XY] = E[X]E[Y] \quad Y^k=1$$

If  $E[XY] = 0$ , then  $X, Y$  are orthogonal.

**28.2 Central Moment**

$$E[(X - E[X])^j \cdot (Y - E[Y])^k]$$

**28.2.1 Variance**

$$\text{VAR}(X) = E[(X - E[X])^2] = E[(X - E[X])^2]$$

**29 Covariance**

$$\text{COV}(X, Y) = E[(X - E[X])(Y - E[Y])]$$

$$= E[(X - E[X])^1 \cdot (Y - E[Y])^1] = E[XY] - E[X]E[Y]$$

If  $E[XY] = 0$  and/or  $E[Y] = 0$ , then

$$\text{COV}(X, Y) = E[XY]$$

**29.1 Correlation Coefficient**

$$\rho_{XY} = \frac{\text{COV}(X, Y)}{\sigma_X \sigma_Y}, \quad -1 \leq \rho_{XY} \leq 1$$

If  $X, Y$  uncorrelated, then

$$\text{COV}(X, Y) = 0, \quad E[XY] = E[X]E[Y], \quad \rho_{XY} = 0$$

If  $X, Y$  independent, then they are uncorrelated.

**29.2 Covariance Properties**

$$\text{COV}(aX, X) = \text{VAR}(X) \quad \text{COV}(X, Y) = \text{COV}(Y, X)$$

$$\text{COV}(aX, Y) = a \text{COV}(X, Y)$$

$$\text{COV}(X + c, Y) = \text{COV}(X, Y)$$

$$\text{COV}(X + Y, Z) = \text{COV}(X, Z) + \text{COV}(Y, Z)$$

**30 Conditional Probabilities**

**30.1 Case 1: X, Y Discrete - Conditional PMF**

$$p_Y(y|x) = P[Y = y|X = x] = \frac{P[XY = x, Y = y]}{P[X = x]}$$

**30.2 Case 2: X Discrete, Y continuous - Conditional PDF**

$$f_Y(y|x_k) = \frac{P[Y \leq y, X = x_k]}{P[X = x_k]} = P[X = x_k > 0]$$

$$f_Y(y|x_k) = \frac{d}{dy} F_Y(y|x_k)$$

**30.3 Case 3: X, Y continuous - Conditional PDF**

$$f_Y(y|x) = \frac{d}{dy} F_Y(y|x) = \frac{f_{XY}(x, y)}{f_X(x)}$$

**30.4 Bayes Rule**

$$P[Y \in A|X = x] = \int_{y \in A} f_Y(y|x) dy$$

$$P[Y \in A|X = x] = \int_{y \in A} f_Y(y|x) dy$$

**31 Conditional Expectation**

**31.1 XY Discrete**

$$E[Y|X] = \sum_{y_k} p_Y(y_k|x)$$

**31.2 XY Continuous**

$$E[Y|X] = \int_{-\infty}^{\infty} y f_Y(y|x) dy$$

**31.3 Law of total Expectation**  
Since  $E[Y|x] = g(X)$ , we define  $E[g(X)]$

$$E[E[Y|X]] = \int_{-\infty}^{\infty} E[Y|x] f_X(x) dx = E[Y]$$

for any function  $h(Y)$ , where  $E[h(Y)] = E[E[h(Y|x)]]$

$$E[Y^k] = E[E[Y^k|x]]$$

**32 Functions of Two RVs**  
Let  $Z = g(X, Y)$  (function of two RVs). Then,

$$F_Z(z) = P[X \in R_Z] = \iint_{(x,y) \in R_Z} f_{XY}(x, y) dx dy$$

$$f_Z(z) = \frac{d}{dz} F_Z(z) = \int_{-\infty}^{\infty} f_{XY}(x, z - x) dx$$

If  $X, Y$  independent, then

$$f_Z(z) = f_X(x) \cdot f_Y(y) = \int_{-\infty}^{\infty} f_X(x) f_Y(z - x) dx$$

**33 Transformations of Two RVs**  
Let  $W = (X, Y)$  and  $Z_1 = g_1(W)$  and  $Z_2 = g_2(W)$

$$F_{Z_1, Z_2}(z_1, z_2) = P[g_1(W) \leq z_1, g_2(W) \leq z_2]$$

$$F_{Z_1, Z_2}(z_1, z_2) = \iint_{W: g_k(W) \leq z_k} f_{XY}(x, y) dx dy$$

**34 Linear Transformations**

$$\begin{bmatrix} W \\ \tilde{W} \end{bmatrix} = \begin{bmatrix} a & b \\ c & d \end{bmatrix} \begin{bmatrix} X \\ Y \end{bmatrix} = A \begin{bmatrix} X \\ Y \end{bmatrix}$$

Assume  $A$  is invertible:

$$\begin{bmatrix} X \\ Y \end{bmatrix} = A^{-1} \begin{bmatrix} W \\ \tilde{W} \end{bmatrix}$$

**34.1 Joint PDF of Linear Transformation**  
Let  $Z = g(X, Y)$ . The vector  $Z$  is:

$$Z = AW \quad Z = \begin{bmatrix} W \\ \tilde{W} \end{bmatrix} \quad W = \begin{bmatrix} X \\ Y \end{bmatrix}$$

The Joint PDF of  $Z$  is

$$f_Z(z) = \frac{f_W(A^{-1}z)}{|A|} \quad |A| = \det \begin{bmatrix} a & b \\ c & d \end{bmatrix}$$

**35 Joint Gaussian RVs**  
The random variables  $X, Y$  are jointly gaussian if:

$$f_{X,Y}(x, y) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho_{XY}^2}} \exp(-A)$$

**35.1 Joint Standard (Normal) Gaussians**  
If  $X \sim N(0, 1), Y \sim N(0, 1)$ , then

$$f_{XY}(x, y) = \frac{1}{2\pi\sqrt{1-\rho_{XY}^2}} \exp(-A)$$

**35.2 Independence (m=0, sigma=1)**  
If  $X, Y$  independent  $\leftrightarrow$

$$\text{COV}(X, Y) = 0 \quad \rho_{XY}(x, y) = 0$$

**35.3 Independence (m=0)**  
If  $X \sim N(0, 1), Y \sim N(0, 1)$ , then

$$f_{XY}(x, y) = \frac{1}{2\pi} \exp\left(-\frac{1}{2}(x^2 + y^2)\right)$$

**35.4 Constant A**  
If  $A$  (exponent of Joint Gaussian) is a constant  $K$

$$f_{XY}(x, y) = \frac{1}{2\pi\sigma^2} \exp\left(-\frac{1}{2\sigma^2}(x^2 + y^2)\right)$$

**35.5 Major Axis**  
If  $X, Y$  not independent, then the principal axes has

$$\theta = \frac{1}{2} \arctan^{-1} \tan\left(\frac{2\rho_{XY}\sigma_1\sigma_2}{\sigma_1^2 - \sigma_2^2}\right)$$

**35.6 Conditional PDF**  
The conditional PDF of  $X$  given  $Y = y$  is

$$f_X(x|y) = \frac{f_{XY}(x, y)}{f_Y(y)} = \frac{1}{2\pi\sigma_1^2\sqrt{1-\rho_{XY}^2}}$$

**35.7 Conditional Expectation**

$$E[(X - m_1)(Y - m_2)|Y] = (y - m_2)E[X - m_1|Y = y]$$

$$= (y - m_2)(\rho_{XY} \frac{\sigma_1}{\sigma_2}(y - m_2)) = \rho_{XY} \frac{\sigma_1}{\sigma_2}(y - m_2)^2$$

**35.8 Covariance**

$$\text{COV}(X, Y) = E[(X - m_1)(Y - m_2)] = E[E[(X - m_1)(Y - m_2)|Y]] = \rho_{XY} \sigma_1 \sigma_2$$

**36 Sum of RVs**  
Let  $X_1, X_2, \dots, X_n$  be a sequence of RVs, with

$$S_n = X_1 + X_2 + \dots + X_n$$

**36.1 Mean and Variance of Sum of RVs**

$$E[X_1 + X_2 + \dots + X_n] = E[X_1] + E[X_2] + \dots + E[X_n]$$

$$\text{VAR}(X_1 + \dots + X_n) = \sum_{k=1}^n \text{VAR}(X_k) + \sum_{j=1}^n \sum_{k=1}^n \text{COV}(X_j, X_k), \quad j \neq k$$

If  $X_1, X_2, \dots, X_n$  independent, then

$$\text{VAR}(X_1 + \dots + X_n) = \text{VAR}(X_1) + \dots + \text{VAR}(X_n)$$

**36.2 PDF of Sums of Independent RVs**  
Let  $X_1, X_2, \dots, X_n$  independent, then

$$\phi_{S_n}(\omega) = E[e^{j\omega S_n}] = E[e^{j\omega(X_1 + X_2 + \dots + X_n)}] = \phi_{X_1}(\omega) \dots \phi_{X_n}(\omega)$$

$$f_{S_n} = \mathcal{F}^{-1}[\phi_{X_1}(\omega) \dots \phi_{X_n}(\omega)]$$

**37 Independent Identically Distributed RVs (iid)**  
If  $X_1, X_2, \dots, X_n$  iid RVs, with

$$E[X_j] = m_X \quad \text{VAR}(X_j) = \sigma_X^2 \quad \text{for } j = 1, \dots, n$$

**37.1 Mean and Variance of iid RVs**

$$E[S_n] = E[X_1] + \dots + E[X_n] = n \cdot m_X$$

$$\text{VAR}(S_n) = n \cdot \text{VAR}(X_j) = n\sigma_X^2$$

**37.2 PDF of iid RVs**

$$\phi_{X_k}(\omega) = \phi_X(\omega), k = 1, \dots, n \leftrightarrow \phi_{S_n}(\omega) = [\phi_X(\omega)]^n$$

$$f_{S_n} = \mathcal{F}^{-1}(\phi_{S_n}(\omega)) = \mathcal{F}^{-1}(\phi_X(\omega)^n)$$

**38 Sample Mean**

$$M_n = \frac{1}{n} \sum_{j=1}^n X_j$$

**38.1 Expected Value and Variance of Sample Mean**

$$E[M_n] = E\left[\frac{1}{n} \sum_{j=1}^n X_j\right] \rightarrow E[M_n] = \frac{1}{n} \sum_{j=1}^n E[X_j]$$

$$\text{VAR}(M_n) = E[(M_n - E[M_n])^2] = \text{VAR}(S_n)/n^2$$

If  $X_1, \dots, X_n$  iid RVs:

$$E[M_n] = m_X \leftrightarrow E[S_n] = n \cdot m_X$$

$$\text{VAR}(S_n) = n\sigma^2 \leftrightarrow \text{VAR}(M_n) = \frac{\sigma^2}{n}$$

**38.2 Sample Mean Chebyshev Bound**

$$P[|Z - E[Z]| \geq \epsilon] \leq \frac{\text{VAR}(Z)}{\epsilon^2}, \quad \epsilon > 0$$

$$P[|M_n - m_X| \geq \epsilon] \leq \frac{\sigma^2}{n\epsilon^2}$$

$$P[|M_n - m_X| < \epsilon] \geq 1 - \frac{\sigma^2}{n\epsilon^2}$$

**39 Laws of Large Numbers**

Weak Law :  $\lim_{n \rightarrow \infty} P[|M_n - m_X| < \epsilon] = 1$

Strong Law :  $P\left[\lim_{n \rightarrow \infty} M_n = m_X\right] = 1$

**40 Central Limit Theorem**  
Let  $S_n = X_1, X_2, \dots, X_n$  iid RVs

$$\lim_{n \rightarrow \infty} P[Z_n \leq z] = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^z e^{-x^2/2} dx$$

**41 Common Sums**

$$\sum_{k=0}^n a^k = \frac{1-a^{n+1}}{1-a}, a \neq 1$$

$$\sum_{k=0}^n ka^k = \frac{a}{(1-a)^2} [1 - (n+1)a^n + na^{n+1}], a \neq 1$$

$$\sum_{k=0}^n k = \frac{n(n+1)}{2}$$

$$\sum_{k=0}^n k^2 = \frac{(n+1)(2n+1)n}{6}$$

$$\sum_{k=0}^n k^3 = \frac{(n+1)^2 n^2}{4}$$

$$\sum_{k=0}^{\infty} a^k = \frac{1}{1-a}$$

$$\sum_{k=0}^{\infty} ka^k = \frac{a}{(1-a)^2}$$

$$\sum_{k=0}^{\infty} k^2 a^k = \frac{a^2+a}{(1-a)^3}$$

**42 Common Discrete Random Variables**

| RV           | S                               | p_X(k)                                      | E[X]    | VAR[X]                | G_X(z)              |
|--------------|---------------------------------|---|---------|-----------------------|---------------------|
| Uniform      | {1, 2, ..., L}                  | 1/L   | (L+1)/2 | (L^2-1)/12            | z \frac{1-z^L}{1-z} |
| Bernoulli    | {0, 1}                          | (1-p)p                                      | p       | p(1-p)                | q + pz              |
| Binomial     | {0, 1, ..., n}                  | \binom{n}{k} p^k (1-p)^{n-k}                | np      | np(1-p)               | (q + pz)^n          |
| -ve Binomial | {r, r+1, ...}, r \in \mathbb{R} | \binom{r-1}{k-1} p^{r-k}                    | r/p     | r(1-p)/p^2            | (\frac{pz}{1-pz})^r |
| Geometric v1 | {0, 1, 2, ...}                  | (1-p)^k p                                   | 1/p     | 1/p^2                 | \frac{p}{1-pz}      |
| Geometric v2 | {1, 2, ...}                     | (1-p)^{k-1} p                               | 1/p     | 1/p^2                 | \frac{pz}{1-pz}     |
| Poisson      | {0, 1, 2, ...}                  | \frac{a^k}{k!} e^{-a}, k = 0, 1, ..., a > 0 | a       | a                     | e^{a(z-1)}          |
| Zipf         | {1, 2, ..., L}                  | 1/c_L k^{-c_L}, c_L = \sum_{j=1}^L 1/j      | 1/c_L   | L(L+1)/2c_L - 1/c_L^2 |                     |

**43 Common Continuous Random Variables**

| RV                     | S                  | f_X(x)  | E[X]                        | VAR[X]   | \phi_X(\omega)  | F_X(x)  |
|------------------------|--------------------|---|-----------------------------|--|---|---|
| Uniform                | [a, b]             | 1/(b-a)   | (a+b)/2                     | (b-a)^2/12   | \frac{e^{j\omega b} - e^{j\omega a}}{j\omega(b-a)}            | \frac{x-a}{b-a}, x \in [a, b]   |
| Exponential            | (0, \infty)        | \lambda e^{-\lambda x}  | 1/\lambda                   | 1/\lambda^2  | \frac{1}{1 - j\omega/\lambda}                                 | \phi(x)   |
| Gaussian (Normal)      | (-\infty, \infty)  | \frac{1}{\sigma\sqrt{2\pi}} e^{-(x-\mu)^2/2\sigma^2}                                  | \mu                         | \sigma^2   | e^{j\omega\mu - \sigma^2\omega^2/2}                           |   |
| Gamma                  | (0, \infty)        | \frac{\lambda^\alpha \Gamma(\alpha)^{-1} x^{\alpha-1} e^{-\lambda x}}{\Gamma(\alpha)} | \alpha/\lambda              | \alpha/\lambda^2                                     | (\frac{1}{1-j\omega/\lambda})^\alpha, \alpha > 0, \lambda > 0 |   |
| m-1 Erlang             | (0, \infty)        | \frac{\lambda^m (m-1)! x^{m-1} e^{-\lambda x}}{(m-1)!}                                | m/\lambda                   | m/\lambda^2  | (\frac{1}{1-j\omega/\lambda})^m                               |   |
| \chi^2-Squared (k DoF) | (0, \infty)        | \frac{\lambda^{k/2} e^{-\lambda x/2} x^{k/2-1}}{2^{k/2} \Gamma(k/2)}                  | k/2                         | k/2  | (1 - j\omega/\lambda)^{-k/2}                                  |   |
| Laplacian              | (-\infty, \infty)  | \frac{\alpha}{2} e^{-\alpha x }   | 0                           | \frac{2}{\alpha^2}                                   | \frac{\alpha^2}{\omega^2 + \alpha^2}                          | \int_{-\infty}^x \frac{1}{2} e^{-\alpha t} dt, x \le 0<br>1 - \int_x^{\infty} \frac{1}{2} e^{-\alpha t} dt, x \ge 0 |
| Rayleigh               | [0, \infty)        | \frac{x}{\sigma^2} e^{-x^2/2\sigma^2}   | \sigma\sqrt{\pi}/2          | (2-\pi/2)\sigma^2                                    | \frac{\alpha^2}{1 - e^{-\alpha^2/(2\sigma^2)}}                |   |
| Pareto                 | {x_M, \infty)      | \frac{\alpha x_M^\alpha}{x^{\alpha+1}}, x \ge x_M, \alpha > 0                         | \frac{\alpha x_M}{\alpha-1} | \frac{\alpha x_M^2}{(\alpha-2)(\alpha-1)^2}          | 1 - (\frac{x_M}{x})^\alpha                                    |   |
| Cauchy                 | (-\infty, +\infty) | \frac{1}{\pi} \frac{\sigma}{(x-x_0)^2 + \sigma^2}                                     | x_0                         | \infty   | e^{-\sigma \omega }   |   |
| Beta                   | 0 < x < 1          | \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1} | \frac{\alpha}{\alpha+\beta} | \frac{\alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)} |   |   |

**44 Q(X) for STANDARD Gaussian**

| x   | Q(x)     | Approx.  | x    | Q(x)     | Approx.  |
|-----|----------|----------|------|----------|----------|
| 0   | 5.00E-01 | 5.00E-01 | 2.7  | 3.47E-03 | 3.46E-03 |
| 0.1 | 4.60E-01 | 4.58E-01 | 2.8  | 2.56E-03 | 2.55E-03 |
| 0.2 | 4.21E-01 | 4.17E-01 | 2.9  | 1.87E-03 | 1.86E-03 |
| 0.3 | 3.82E-01 | 3.78E-01 | 3.0  | 1.35E-03 | 1.35E-03 |
| 0.4 | 3.44E-01 | 3.41E-01 | 3.1  | 9.69E-04 | 9.66E-04 |
| 0.5 | 3.09E-01 | 3.05E-01 | 3.2  | 6.88E-04 | 6.86E-04 |
| 0.6 | 2.74E-01 | 2.71E-01 | 3.3  | 4.83E-04 | 4.83E-04 |
| 0.7 | 2.42E-01 | 2.39E-01 | 3.4  | 3.37E-04 | 3.36E-04 |
| 0.8 | 2.12E-01 | 2.09E-01 | 3.5  | 2.33E-04 | 2.32E-04 |
| 0.9 | 1.83E-01 | 1.82E-01 | 3.6  | 1.59E-04 | 1.59E-04 |
| 1.0 | 1.59E-01 | 1.57E-01 | 3.7  | 1.08E-04 | 1.08E-04 |
| 1.1 | 1.36E-01 | 1.34E-01 | 3.8  | 7.24E-05 | 7.23E-05 |
| 1.2 | 1.15E-01 | 1.14E-01 | 3.9  | 4.81E-05 | 4.81E-05 |
| 1.3 | 9.68E-02 | 9.60E-02 | 4.0  | 3.17E-05 | 3.16E-05 |
| 1.4 | 8.08E-02 | 8.01E-02 | 4.5  | 3.40E-06 | 3.40E-06 |
| 1.5 | 6.68E-02 | 6.63E-02 | 5.0  | 5.51E-07 | 5.51E-07 |
| 1.6 | 5.48E-02 | 5.44E-02 | 5.5  | 1.90E-08 | 1.90E-08 |
| 1.7 | 4.46E-02 | 4.43E-02 | 6.0  | 9.87E-10 | 9.86E-10 |
| 1.8 | 3.59E-02 | 3.57E-02 | 6.5  | 4.02E-11 | 4.02E-11 |
| 1.9 | 2.87E-02 | 2.86E-02 | 7.0  | 1.28E-12 | 1.28E-12 |
| 2.0 | 2.28E-02 | 2.26E-02 | 7.5  | 3.19E-14 | 3.19E-14 |
| 2.1 | 1.79E-02 | 1.78E-02 | 8.0  | 6.49E-16 | 6.49E-16 |
| 2.2 | 1.39E-02 | 1.39E-02 | 8.5  | 9.48E-18 | 9.48E-18 |
| 2.3 | 1.07E-02 | 1.07E-02 | 9.0  | 1.13E-19 | 1.13E-19 |
| 2.4 | 8.20E-03 | 8.17E-03 | 9.5  | 1.05E-21 | 1.05E-21 |
| 2.5 | 6.21E-03 | 6.19E-03 | 10.0 | 7.62E-24 | 7.62E-24 |
| 2.6 | 4.66E-03 | 4.65E-03 |      |          |          |